On Constrained Optimization of the Klimov Network and Related Markov Decision Processes

by A.M. Makowski and A. Shwartz

TECHNICAL RESEARCH REPORT



Supported by the
National Science Foundation
Engineering Research Center
Program (NSFD CD 8803012),
the University of Maryland,
Harvard University,
and Industry

April 1991

ON CONSTRAINED OPTIMIZATION OF THE KLIMOV NETWORK AND RELATED MARKOV DECISION PROCESSES *

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ABSTRACT

We solve a constrained version of the server allocation problem for the Klimov network and establish that the optimal constrained schedule is obtained by randomizing between two fixed priority schemes. This generalizes work of Nain and Ross in the context of the competing queue problem, and also covers the discounted cost case.

In order to establish these results we develop a general framework for optimization under a single constraint in the presence of index-like policies. This methodology is in principle of wider applicability.

Keywords: Klimov queueing network, Constrained Optimization, Bandit Processes.

^{*} Research supported partially through NSF Grant NSF ECS-83-51836 and partially through United States—Israel Binational Science Foundation Grant BSF85-00306.

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1. INTRODUCTION

Consider the discrete-time system of K competing queues with a single Bernoulli server as described in [5, 8]. For one-step costs which are linear in the queue sizes, it is well known [4, 5] that there exists an optimal policy which is of the strict priority type, and this under several cost criteria, including the discounted and average cost criteria. In these situations, the search for optimal policies simply reduces to the computation of a few parameters. Renewed interest over the past decade in stochastic control problems with constraints led Nain and Ross [19] to investigate a simple scheduling problem for the discrete-time system of competing queues: Let $J_c(\pi)$ and $J_d(\pi)$ be two cost functions, associated with the one-step cost functions c and d, when the system is operated under the policy π . A single constraint optimization problem can then be defined as

$$(\mathbf{P}_{\mathbf{V}})$$
: Minimize $J_c(\pi)$ subject to the constraint $J_d(\pi) \leq V$

for some scalar V. When both costs c and d are linear in the queue sizes, under the average cost criterion, Nain and Ross [19] obtained the following optimality result: There exist two fixed priority policies, say \underline{g} and \overline{g} , and a constant η^* in [0,1] so that, at every step, it is optimal to flip a coin with probability η^* for heads, and to use \underline{g} (resp. \overline{g}) if a head (resp. tail) is observed. The optimal randomization bias η^* is selected so as to saturate the constraint.

In view of such results, it is quite natural to inquire whether this structural result for the optimal constrained policy can be extended, say to cover

- (i) the situation where the discounted or the finite-time cost criteria would be used;
- (ii) the scheduling problem associated with a natural extension of the competing queue problem, namely the so-called Klimov system [14], where upon service completion the customer may either be routed to one of the other queues or leave the system.

More generally, one can certainly wonder as to the conditions under which the solution to a constrained MDP does exhibit such a structure. The interest in establishing such a structural optimality result should be quite clear. Indeed, once established, the search for optimal policies is reduced to the identification of the two policies and to the computation of the randomization bias.

We answer (i)-(ii) in the affirmative in Section 5. In the process, we develop in Section 3 a more general methodology which applies to systems with "index-like" optimal policies. This is embedded in the only "structural" assumption (A1), whereby an optimal policy for the *unconstrained* problem with cost $c + \theta d$ can be found from a fixed, finite set of policies, where this set is independent of $\theta \in [0,1]$. In Section 4 we show that the technical conditions apply to many MDPs, with

finite, discounted and average cost criteria. In Section 5 we establish the equivalence between the discounted Klimov system and open (or arm-acquiring) bandit problems. This establishes the structural result for the Klimov system under the discounted cost criterion, and for all bandits problems under mild conditions. Under slightly stronger assumptions, the structural result also holds for the Klimov system under the average-cost criterion.

Of course, optimization problems under constraints are not new. The first such problem was solved for the finite-horizon cost by Derman and Klein [9]. When $J(\pi)$ is the average cost criterion, the existence of optimal stationary policies under multiple constraints was established by Derman and Veinott [10], and by Hordijk and Kallenberg [13], both for finite state space S and action space U, including the multi-class case. Under a single class assumption and for a single constraint, the existence of optimal stationary policies which are randomized at a single state was proved by Beutler and Ross [6] for finite S and compact U, and by Sennott [22] for countable S and compact S and compact S and action when S is countable and S is countable and S is countable and S is countable and S is countable S and S

Unfortunately, except for the finite case and the specific example in [1], there are no efficient methods for the computation of optimal policies. The results alluded to in the previous paragraph establish the existence of an optimal stationary policy which randomizes between some stationary deterministic policies. However, except for the finite case, the search for the two policies to randomize is over all stationary deterministic policies. Our methodology provides conditions under which this search can be restricted to a finite set of policies.

A few words on the notation and conventions used in this paper: For any set E endowed with a topology, measurability is always taken to mean Borel measurability and the corresponding Borel σ -field, i.e., the smallest σ -field on E generated by the open sets of the topology, is denoted by $\mathcal{B}(E)$. Unless otherwise stated, $\lim_n \lim_n$ and $\overline{\lim}_n$ are taken with n going to infinity. Moreover, the infimum over an empty set is taken to be ∞ by convention.

2. THE GENERAL MODEL

To set up the discussion, we start with a MDP (S, U, P) as defined in the literature [20, 21, 25]. The state space S and the action space U are assumed to be Polish spaces; the one-step transition mechanism P is defined through the family (Q(x, u; dy)) of measurable transition kernels. The state process $\{X_t, t = 0, 1, \ldots\}$ and the control process $\{U_t, t = 0, 1, \ldots\}$ are defined on some measurable space (Ω, \mathcal{F}) (which for sake of concreteness is taken to be the canonical space $S \times (U \times S)^{\infty}$ equipped with the product topology). The feedback information available to the decision-maker is encoded through the random variables (rvs) $\{H_t, t = 0, 1, \ldots\}$ defined by $H_0 \triangleq X_0$ and by $H_t \triangleq (X_0, U_0, X_1, \ldots, U_{t-1}, X_t)$ for all $t = 1, 2, \ldots$ so that the rvs X_t, U_t and H_t take values in S, U and $H_t \triangleq S \times (U \times S)^t$, respectively; we also define the information σ -field \mathcal{F}_t by $\mathcal{F}_t = \sigma\{H_t\}$.

The space of probability measures on $\mathcal{B}(U)$ is denoted by $\mathbb{M}(U)$. An admissible control policy π is defined as any collection $\{\pi_t,\ t=0,1,\ldots\}$ of mappings $\pi_t:\mathbb{H}_t\to\mathbb{M}(U)$ such that for all $t=0,1,\ldots$ and every Borel subset B of U, the mapping $\mathbb{H}_t\to[0,1]:h_t\to\pi_t(h_t;B)$ is Borel measurable. The collection of all such admissible policies is denoted by \mathcal{P} .

Let μ be a given probability measure on $\mathcal{B}(S)$. The definition of the MDP (S, U, P) then postulates the existence of a collection of probability measures $\{\mathbf{P}^{\pi}, \pi \in \mathcal{P}\}$ on \mathcal{F} such that conditions (2.1)–(2.2) below are satisfied: For every admissible policy π in \mathcal{P} , the probability measure \mathbf{P}^{π} is constructed so that under \mathbf{P}^{π} , the rv X_0 has probability distribution μ , the control actions are selected according to

$$\mathbf{P}^{\pi}[U_t \in B | \mathcal{F}_t] = \pi_t(H_t; B), \quad B \in \mathcal{B}(U)$$
 $t = 0, 1, \dots (2.1)$

and the state transitions are realized according to

$$\mathbf{P}^{\pi}[X_{t+1} \in A | \mathcal{F}_t \vee \sigma\{U_t\}] = Q(X_t, U_t; A), \quad A \in \mathcal{B}(S).$$
 $t = 0, 1, \dots (2.2)$

The expectation operator associated with \mathbf{P}^{π} is denoted by \mathbf{E}^{π} .

Following standard usage, a policy π in \mathcal{P} is said to be a Markov policy if there exists a family $\{g_t, t=0,1,\ldots\}$ of Borel mappings $g_t:S\to \mathbb{M}(U)$ such that $\pi_t(\cdot;H_t)=g_t(\cdot;X_t)$ \mathbf{P}^{π} -a.s. for all $t=0,1,\ldots$ In the event the mappings $\{g_t, t=0,1,\ldots\}$ are all identical to a given mapping $g:S\to \mathbb{M}(U)$, the Markov policy is termed stationary and is identified with the mapping g itself.

3. A GENERAL CONSTRAINED MDP

We interpret any Borel mapping $c: S \times U \to \mathbb{R}$ as a one-step cost function. In order to avoid unnecessary technicalities we always assume c to be bounded below. In fact, as will be apparent from the discussion, there is no loss of generality in assuming $c \geq 0$, as we do from now on. For any policy π in \mathcal{P} , we define $J_c(\pi)$ as the total cost (associated with c) for operating the system under policy π . Several choices are possible and include the long-run average cost

$$J_c(\pi) \stackrel{\triangle}{=} \overline{\lim}_t \mathbf{E}^{\pi} \left[\frac{1}{t+1} \sum_{s=0}^t c(X_s, U_s) \right], \tag{3.1}$$

and the infinite-horizon β -discounted cost

$$J_c(\pi) \stackrel{\triangle}{=} \mathbf{E}^{\pi} \left[\sum_{s=0}^{\infty} \beta^s c(X_s, U_s) \right], \quad 0 < \beta < 1.$$
 (3.2)

The definitions (3.1)–(3.2) are all well posed under the non-negativity assumption on c.

Now, we consider two Borel mappings $c, d: S \times U \to \mathbb{R}_+$ and for some scalar V, we set

$$\mathcal{P}_V := \{ \pi \in \mathcal{P} : J_d(\pi) \le V \}. \tag{3.3}$$

The corresponding *constrained* optimization problem $(\mathbf{P}_{\mathbf{V}})$ is now formulated as

$$(\mathbf{P}_{\mathbf{V}})$$
: Minimize $J_c(\cdot)$ over \mathcal{P}_V .

Implicit in this formulation is the fact that the cost criteria $J_c(\cdot)$ and $J_d(\cdot)$ are of the same type.

For every θ in [0,1], we define the mapping $c_{\theta}: S \times U \to \mathbb{R}_+$ by

$$c_{\theta}(x,u) \stackrel{\triangle}{=} \theta c(x,u) + (1-\theta)d(x,u), \quad x \in S, u \in U.$$
(3.4)

We alleviate somewhat the notation by using $J_{\theta}(\pi)$ to denote the total cost associated with c_{θ} under policy π , whence $J_{\theta}(\pi) = J_{c}(\pi)$ for $\theta = 1$ and $J_{\theta}(\pi) = J_{d}(\pi)$ for $\theta = 0$. The discussion is given under the following general assumptions (A1), where

(A1) There exists a finite number of Markov stationary policies g_1, \ldots, g_L such that

(A1.a) The condition

$$\inf_{\pi \in \mathcal{P}} J_{\theta}(\pi) = \min_{1 \le \ell \le L} J_{\theta}(g_{\ell}) \stackrel{\triangle}{=} J^{*}(\theta), \quad \theta \in [0, 1]$$
(3.5)

holds true; and

(A1.b) For each $\ell = 1, ..., L$, the mapping $\theta \to J_{\theta}(g_{\ell})$ is continuous on [0, 1].

It is plain that under (A1), the mapping $\theta \to J^*(\theta)$ is continuous on [0, 1]. As in [1], the Lagrangian problem is defined as the problem of minimizing $J_{\theta}(\cdot)$ over the unconstrained set of policies \mathcal{P} . We define

$$N(\theta) \stackrel{\triangle}{=} \left\{ \ell \in \{1, \dots, L\} : J_{\theta}(g_{\ell}) = J^{*}(\theta) \right\}, \quad \theta \in [0, 1].$$

$$(3.6)$$

Using (A1) we readily obtain the following properties: For each θ in [0,1], the index set $N(\theta)$ is always non-empty by virtue of (A1.a), and (A1.b) implies

$$\lim_{\tilde{\theta}\uparrow\theta} J_{\tilde{\theta}}(g_{\ell}) = \lim_{\tilde{\theta}\downarrow\theta} J_{\tilde{\theta}}(g_{\ell}) = J^{*}(\theta) , \quad \ell \in N(\theta) . \tag{3.7}$$

Furthermore, if $N(\theta)$ reduces to a singleton, then $N(\theta) = N(\tilde{\theta})$ in some open neighborhood of θ .

To proceed, we set

$$n(\theta) \stackrel{\triangle}{=} \min \left\{ n \in N(\theta) : J_d(g_n) = \min_{\ell \in N(\theta)} J_d(g_\ell) \right\}, \quad \theta \in [0, 1].$$
 (3.8)

If $J_0(g_{n(0)}) = J_d(g_{n(0)}) > V$, then the problem $(\mathbf{P}_{\mathbf{V}})$ is not feasible and therefore possesses no solution. Assuming feasibility from now on, we set

$$\theta^* \stackrel{\triangle}{=} \sup\{\theta \in [0,1] : J_d(g_{n(\theta)}) \le V\}. \tag{3.9}$$

If $\theta^* = 0$, then necessarily $J_d(g_{n(0)}) \leq V$, but we may have to entertain the possibility that

$$\min\{J_c(g_\ell): 1 \le \ell \le L, J_d(g_\ell) \le V\} > \inf_{\pi \in \mathcal{P}_V} \{J_c(\pi): J_d(\pi) \le V\}$$

since the Lagrangian problem may not provide enough information.

If $\theta^* = 1$, then $(\mathbf{P_V})$ has a solution: Indeed, let $\theta_i \uparrow 1$ in (0,1] so that $J_d(g_{n(\theta_i)}) \leq V$ for all i = 1, 2, ... by the definition of θ^* . A converging subsequence, say $\theta_j \uparrow 1$, can always be selected so that $n(\theta_j) \to n^*$ for some n^* in $\{1, ..., L\}$. In fact, we can assert $n(\theta_j) = n^*$ whenever $j \geq j^*$ for some j^* . It is plain that n^* is an element of $N(\theta_j)$ for $j \geq j^*$, whence $J_{\theta_j}(g_{n^*}) = J^*(\theta_j)$. The continuity of $\theta \to J^*(\theta)$ implies that n^* is an element of N(1), and since $J_d(g_{n^*}) \leq V$, we conclude that the policy g_{n^*} solves $(\mathbf{P_V})$.

From now on, assume $0 < \theta^* < 1$. Let $\theta_i \downarrow \theta^*$ in (0,1) and denote by \overline{n} an accumulation point of the sequence $\{n(\theta_i), i=1,2,\ldots\}$. Similarly, let $\theta_j \uparrow \theta^*$ in (0,1) such that $J_d(g_{n(\theta_j)}) \leq V$

and denote by \underline{n} an accumulation point of $\{n(\theta_j), j = 1, 2, \ldots\}$. Again, we have $n(\theta_i) = \overline{n}$ and $n(\theta_j) = \underline{n}$ for all i and j large enough. By (A1.b), we see that both \overline{n} and \underline{n} are elements of $N(\theta^*)$, so that the equalities

$$J_{\theta^*}(g_n) = J_{\theta^*}(g_{\overline{n}}) = J^*(\theta^*) \tag{3.10}$$

must hold. Moreover, it is plain that

$$J_d(g_n) \le V \le J_d(g_{\overline{n}}). \tag{3.11}$$

The first inequality follows by construction and (A1.b), whereas the second inequality results from the construction and (3.8)–(3.9).

Next, we define the policies $\underline{g}, \overline{g}$ and $\{g^{\eta}, 0 \leq \eta \leq 1\}$, as the Markov stationary policies given by

$$g \stackrel{\triangle}{=} g_{\underline{n}}, \quad \overline{g} \stackrel{\triangle}{=} g_{\overline{n}} \tag{3.12}$$

and

$$g^{\eta} \stackrel{\triangle}{=} \eta g + (1 - \eta)\overline{g}, \quad \eta \in [0, 1].$$
 (3.13)

Then g^{η} is the simple randomization between the two policies \underline{g} and \overline{g} with randomization bias η . The identities (3.10)–(3.11) now take the form

$$J_{\theta^*}(\underline{g}) = J_{\theta^*}(\overline{g}) = J^*(\theta^*)$$
(3.14)

and

$$J_d(g) \le V \le J_d(\overline{g}). \tag{3.15}$$

At this point, we can introduce the condition (A2).

(A2) The mapping $\eta \to J_d(g^{\eta})$ is continuous on [0,1].

Lemma 1. Under (A1)-(A2), the equation

$$J_d(q^{\eta}) = V, \quad \eta \in [0, 1]$$
 (3.16)

has a solution η^* .

Proof. This is immediate from the fact that the mapping $\eta \to J_d(g^{\eta})$ is continuous on [0,1] and from the inequality (3.15) which can written as $J_d(g^1) \le V \le J_d(g^0)$.

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We further assume that conditions (A3)-(A5) are enforced, where

(A3) The equality

$$J_{\theta^*}(g^{\eta}) = J_{\theta^*}(g) , \quad \eta \in [0, 1]$$
 (3.17)

holds;

(A4) The equality

$$J_{\theta^*}(g^{\eta^*}) = \theta^* J_c(g^{\eta^*}) + (1 - \theta^*) J_d(g^{\eta^*})$$
(3.18)

holds; and

(A5) For every admissible policy π in \mathcal{P} , the inequality

$$J_{\theta^*}(\pi) \le \theta^* J_c(\pi) + (1 - \theta^*) J_d(\pi) \tag{3.19}$$

holds.

Theorem 2. Under (A1)-(A5), the policy g^{η^*} (where η^* is a solution of (3.16)) solves the constrained problem ($\mathbf{P}_{\mathbf{V}}$) provided $\theta^* > 0$.

Proof. We first note that

$$J^*(\theta^*) = J_{\theta^*}(g^{\eta^*}) \tag{3.20}$$

$$= \theta^* J_c(g^{\eta^*}) + (1 - \theta^*) J_d(g^{\eta^*})$$
(3.21)

where (3.20) follows from (3.14) and (A3), whereas (3.21) is validated by (A4). On the other hand, we have

$$J_{\theta^*}(\pi) \ge J^*(\theta^*), \quad \pi \in \mathcal{P}$$
 (3.22)

by virtue of (A1.a), and

$$J_{\theta^*}(\pi) \le \theta^* J_c(\pi) + (1 - \theta^*) J_d(\pi), \quad \pi \in \mathcal{P}$$
(3.23)

by invoking (A5). By Lemma 1, the policy g^{η^*} is an element of \mathcal{P}_V since $J_d(g^{\eta^*}) = V$ by construction, and upon combining (3.20)–(3.23), we get

$$\theta^* J_c(\pi) + (1 - \theta^*) J_d(\pi) \ge J_{\theta^*}(\pi) \ge \theta^* J_c(g^{\eta^*}) + (1 - \theta^*) V, \quad \pi \in \mathcal{P}$$
 (3.24)

It is now plain from (3.24) that

$$\theta^* J_c(g^{\eta^*}) \le \theta^* J_c(\pi), \quad \pi \in \mathcal{P}_V$$
 (3.25)

and the result follows since $\theta^* > 0$.

Theorem 2 and its proof remain unchanged if (A2) is replaced by the conclusion of Lemma 1, namely that

(A2bis) There exists a solution to equation (3.16),

and if, in addition, (3.17) is assumed to hold only for $\eta = \eta^*$. However, (A2)-(A3) seem more natural and hold under weak conditions, as established in Section 4. Moreover, usually, η^* is not known and therefore (A2)-(A4) are verified by establishing the conditions for all η in [0,1].

We conclude this section by noting that the Markovian properties and the specific structure of the cost criterion are not used in the proof of Theorem 2, in that the discussion applies to any optimization problem which satisfies conditions (A1)-(A5). The only point which requires special care is the construction of an "interpolated" policy (3.13).

In particular, consider the finite-horizon β -discounted cost

$$J_c(\pi) \stackrel{\triangle}{=} \mathbf{E}^{\pi} \left[\sum_{s=0}^{T} \beta^s c(X_s, U_s) \right], \quad 0 < \beta \le 1, \ T = 1, 2, \dots$$
 (3.26)

The derivation of Lemma 1 and Theorem 2 holds verbatim, provided (A1) holds with the word "stationary" omitted. Since the identification of a policy g with a function $g: S \to \mathbb{M}(U)$ does not hold, (3.13) is interpreted naturally as

$$g_t^{\eta} \stackrel{\triangle}{=} \eta \underline{g}_t + (1 - \eta) \overline{g}_t, \quad \eta \in [0, 1] .$$
 $t = 0, 1, \dots (3.27)$

4. THE ASSUMPTIONS

In this section we discuss the assumptions (A1)-(A5); we give concrete and verifiable conditions for several cost criteria. A specific model is analyzed in Section 5.

We focus on three cost criteria. The infinite-horizon β -discounted cost (3.2) and the long-run average cost criterion (3.1) are the most common criteria in applications. The are often approximated by, or serve as an approximation for the finite-time β -discounted cost criterion (3.26). The discussion and methods apply, *mutatis mutandis*, to other situations as well. However, for the sake of brevity, we shall not elaborate in that direction.

The finite-time cost criterion – Condition (A2) holds if the costs are bounded since then the costs are polynomial in η . More generally, the same argument establishes (A2) if the costs are merely bounded from below (or from above).

Assumption (A3) holds if (3.5) is valid for all initial conditions, since then a backward-induction argument proves that for any η in [0,1], g^{η} is optimal for the Lagrangian problems. Finally, (A4)-(A5) are always valid since under the non-negativity assumption on c and d, the equality

$$J_{\theta}(\pi) = \theta J_{c}(\pi) + (1 - \theta)J_{d}(\pi), \quad \theta \in [0, 1]$$
 (4.1)

holds for every admissible policy π in \mathcal{P} . Condition (A1.b) immediately follows.

The discounted cost criterion – Condition (A2) holds if the costs are bounded since then the total discounted cost can be approximated by a finite number of terms in (3.2) uniformly in η , and the argument for the finite case applies. More generally, under the same conditions as for the finite cost, the same argument applies provided a finite approximation is valid. This is the case if the tail of the infinite sum is bounded for η in [0,1]. This condition holds for all but the most pathological systems.

Assumption (A3) holds under rather weak conditions. For example, suppose the action space is compact and the costs bounded above. Assume further that for each x in S, the mappings $u \to c(x,u)$ and $u \to d(x,u)$ are lower-semi continuous and that the transition kernel $Q(x;\cdot;dy)$ is weakly continuous (that is, whenever $c:S\to\mathbb{R}$ is bounded and continuous, the mapping $u\to \int c(y)\ dQ(x,u,dy)$ is continuous on U for each x in S). Then any policy with actions in the optimal set (determined through the dynamic programming equation) is optimal for the Lagrangian problem [21]. This implies that (3.17) holds whenever (3.5) is valid for each initial condition. Note that in this case boundedness from above replaces boundedness from below.

Finally, (A4)-(A5) always hold since, as in the finite case, (4.1) holds, and condition (A1.b) immediately follows.

The long-run average cost criterion – Condition (A2) was established when the state space S is finite in [16], and for the queueing system discussed in the next section [18]. A general method for verifying (A2) is available in [23]. In particular, this condition holds whenever the Markov chain is ergodic under both \underline{g} and \overline{g} , provided the costs are integrable under the resulting invariant measures [16].

Condition (A3) can be established using dynamic programming arguments, as in the case of the discounted cost, although the requisite conditions are more stringent [21, 25]. For some systems (such as the one described in Section 5), (A3) can be established by direct arguments [5, 18].

Finally, we observe that for every admissible policy π in \mathcal{P} , the inequalities

$$J_{\theta}(\pi) = \overline{\lim}_{t} E^{\pi} \left[\frac{1}{t+1} \sum_{s=0}^{t} c_{\theta}(X_{s}, U_{s}) \right]$$

$$= \overline{\lim}_{t} \left\{ \theta E^{\pi} \left[\frac{1}{t+1} \sum_{s=0}^{t} c(X_{s}, U_{s}) \right] + (1-\theta) E^{\pi} \left[\frac{1}{t+1} \sum_{s=0}^{t} d(X_{s}, U_{s}) \right] \right\}$$

$$\leq \theta \overline{\lim}_{t} E^{\pi} \left[\frac{1}{t+1} \sum_{s=0}^{t} c(X_{s}, U_{s}) \right] + (1-\theta) \overline{\lim}_{t} E^{\pi} \left[\frac{1}{t+1} \sum_{s=0}^{t} d(X_{s}, U_{s}) \right]$$

$$= \theta J_{c}(\pi) + (1-\theta) J_{d}(\pi), \quad \theta \in [0, 1]$$

$$(4.2)$$

always hold, so that condition (A5) is always satisfied. The validity of (A4) is more delicate to establish. In [23], the authors give conditions under which the long-run average cost criterion (3.1) is obtained as a limit under stationary policies. Under these conditions, (A4) holds, and (A1.b) follows.

5. BANDITS AND QUEUES

The purpose of this section is to show the equivalence between the discrete-time Klimov problem [14, 17] and arm-acquiring bandit processes [24]. Continuous-time versions of this result are discussed in [15, 25]. Since both systems were discussed in detail elsewhere, we shall give only short, informal descriptions. Throughout this section, the rv ξ and the i.i.d. sequence $\{A(t), t = 0, 1, ...\}$ taking their values in \mathbb{N}^K are held fixed. We introduce the finiteness assumption

(F)
$$\mathbf{E}[\xi_k] < \infty \text{ and } \mathbf{E}[A_k(t)] \stackrel{\triangle}{=} \lambda_k < \infty, \quad k = 1, 2, \dots, K.$$

Arm-acquiring bandits.

The formulation is given in the terminology of queueing systems in order to facilitate the comparison: Customers of type 1, 2, ..., N arrive into the system; a customer of type n can be in one of the states $\{1, 2, ..., S_n\}$. It is convenient to lump together customers sharing both type and state [24]; we shall say that a customer of type n in state s, $1 \le s \le S_n$, resides in queue k, where

$$k = \sum_{j=1}^{n-1} S_j + s \tag{5.1}$$

and where $K = \sum_{n=1}^{N} S_n$. With this convention, the number of customers initially in the system is ξ , and new customers arrive to the queues according to the arrival process $\{A(t), t=0,1,\ldots\}$. At most one customer can be given service attention at a time. If a customer from queue k is served in time slot t, then at the end of the slot, with probability $p_{k\ell}$ this customer moves to queue ℓ , $1 \le k, \ell \le K$. All other customers do not change state—in other words, they remain at their queues. The routing rvs are assumed to form an i.i.d. sequence. It is clear that the vector x in \mathbb{N}^K , where x_k is the number of customers in queue k, serves as a state for this MDP provided arrival, service completion and routing processes are mutually independent. This together with the assumption on the routing mechanism implies that at each queue, the events that a customer leaves the system can be modeled by i.i.d. Bernoulli rvs with queue-dependent parameter. The action u=k is interpreted as service at queue k, u=0 as idle server, with the provision that $x_k=0$ implies $u \ne k$, $k=1,2,\ldots,K$. If a customer in queue k is served, then reward r(k) is incurred. The reward to be maximized is of the discounted type (3.2), and takes the form

$$J_r(\pi; x) \stackrel{\triangle}{=} \mathbf{E}^{\pi} \left[\sum_{s=0}^{\infty} \beta^s r(U_s) \right] , \quad 0 < \beta < 1$$
 (5.2)

which is well defined since r is bounded.

The classical description of the arm-acquiring bandits requires $\sum_{\ell} p_{k\ell} = 1$ for each $1 \le k \le K$. However, this restriction is a purely semantic one since the effect of departures from the system can always be captured through the introduction of an absorbing queue with small (negative) reward for service, so that it is never served.

The discrete-time Klimov problem

Customers of type 1, 2, ..., K arrive to their respective queues according to the arrival process $\{A(t), t = 0, 1, ...\}$. The number of customers present at time 0 is given by ξ . The server can attend at most one queue at a time. If the server attends a non-empty queue, say queue k, $1 \le k \le K$, during time slot t, then at the end of the slot the following sequence of events takes place:

One customer leaves that queue with probability μ_k and, with probability $1 - \mu_k$ no customer leaves that queue;

If a customer has left queue k, then with probability $\tilde{p}_{k\ell}$ it joins queue ℓ , $1 \leq \ell \leq K$, and it leaves the system with probability $1 - \sum_{\ell=1}^{K} \tilde{p}_{k\ell}$.

For $1 \leq k, \ell \leq K$, we set $p_{k\ell} \stackrel{\triangle}{=} \mu_k \tilde{p}_{k\ell}$ for $\ell \neq k$ and $p_{kk} \stackrel{\triangle}{=} 1 - \mu_k (1 - \tilde{p}_{kk})$. Using this transformation, the values of μ_k are henceforth taken to be 1. Then clearly, assuming arrival, service completion and routing processes to be mutually independent, the dynamics of this system are equivalent to the dynamics of the corresponding arm-acquiring bandit system.

The state of this system is again the vector x in \mathbb{N}^K where x_k denotes the number of customers in queue $k, 1 \leq k \leq K$. The cost for the Klimov problem is defined by

$$c(x, u) = c(x) \stackrel{\triangle}{=} \sum_{k=1}^{K} c_k x_k, \qquad x \in \mathbb{N}^K, \ u = 0, 1, \dots, K.$$

for some constants c_1, \ldots, c_K (which are usually assumed non-negative). The objective is to *minimize* the discounted cost associated with this one-step cost, viz.

$$J_c(\pi) \stackrel{\triangle}{=} \mathbf{E}^{\pi} \left[\sum_{s=0}^{\infty} \beta^s c(X_s) \right], \quad \pi \in \mathcal{P}.$$
 (5.3)

Following the cost-transformation technique of [4, 5] it is straightforward to derive the identity

$$J_c(\pi) = \frac{\mathbf{E}\,c(\xi)}{1-\beta} + \frac{\beta}{(1-\beta)^2}c(\lambda) - \frac{\beta}{1-\beta}J_{\tilde{c}}(\pi) \;, \quad \pi \in \mathcal{P}.$$
 (5.4)

where the one-step cost \tilde{c} is defined by

$$\tilde{c}(x,u) \stackrel{\triangle}{=} \sum_{k=1}^{K} \mathbf{1}[u=k]\tilde{c}_k , \quad \tilde{c}_k \stackrel{\triangle}{=} \left[c_k - \sum_{\ell=1}^{K} p_{k\ell} c_\ell \right] , 1 \le k \le K$$
 (5.5)

where action u is defined as in the bandit problem. As a result, for each fixed β in (0,1), we have

$$\arg\min J_c(\pi) = \arg\max J_{\tilde{c}}(\pi) . \tag{5.6}$$

The cost function \tilde{c} depends only on the queue being served, and so is a legitimate cost function for the bandit problem.

The equivalence result

Theorem 3. Any discrete-time Klimov problem defines an arm-acquiring bandit system with the same dynamics. Under (**F**), they possess the same optimal policies, with costs related by (5.4)–(5.5) (with $r(k) \stackrel{\triangle}{=} \tilde{c}_k$, $1 \le k \le K$). Conversely, any arm-acquiring bandit system defines a Klimov problem with the same dynamics. Moreover, Under (**F**), if the vector $r \stackrel{\triangle}{=} (r(1), r(2), \dots, r(K))'$ is in the range of I - P, then the cost in the Klimov problem can be defined so as to satisfy the transformation (5.4)–(5.5) (with $\tilde{c}_k \stackrel{\triangle}{=} r(k)$, $1 \le k \le K$) and consequently, the same policies are optimal for both systems.

The proof follows from the preceding discussion, upon observing that if r is in the range of I-P then there is a one-to-one mapping between (c_1,\ldots,c_K) and $(\tilde{c}_1,\ldots,\tilde{c}_K)$.

Constrained Optimization

The best-known class of problems for which the hypotheses (A1)-(A5) hold is the class of armacquiring (or open) bandit processes [24] described above. For consistency with the notation of Section 3, we let c and d still denote the two cost functions (although in this case they are independent of x).

Lemma 4. For the arm-acquiring bandit problem under the discounted cost criterion, conditions (A1)-(A5) hold.

Proof. It is well known [24] that the optimal policy for this system possesses an index-structure. Thus an optimal policy (for any $0 \le \theta \le 1$) chooses only which queue to serve. Therefore such a policy is uniquely determined by an ordering of the queues, where a queue is served only if queues with higher priority are empty. Since there is a finite number K! of such policies, (A1.a) follows. Since the costs are bounded and the action space is discrete, the argument in Section 4 now establishes the result.

We call the Klimov problem stable if $\rho \triangleq \lambda'(I-P)e < 1$ (where e is the element of \mathbb{N}^K with all unit entries, i.e., $e = (1, \dots, 1)'$). A policy is called non-idling if $x_k = 0$ implies $u \neq k$.

Lemma 5. Consider the average-cost case. Assume (**F**) and that the Klimov problem is stable. Moreover, let $c_k \geq 0$, $1 \leq k \leq K$. (i) If $\{g_\ell, \ell = 1, 2, ..., L\}$ is a collection of stationary non-idling policies, then (A1.b) and (A2)-(A5) hold. (ii) If P is diagonal then (A1.a) holds, where $\{g_\ell, \ell = 1, 2, ..., L\}$ is a collection of strict priority policies.

Proof. Under the conditions in (i), Makowski and Shwartz [17, 23] establish (A2), whereas (A4) follows from [23]. As discussed in Section 4, (A5) holds, and (A1.b) follows from (A4). Finally, under the regularity conditions established in [17], standard dynamic programming techniques yield (A3). Part (ii) is established in [4, 5].

When P is diagonal, Theorem 2 now implies the existence of an optimal policy which randomizes between two strict priority policies, and we recover the results of [19]. In general, if we strengthen (F) to require finite second moments, then [17, 18] establish that for every stationary non-idling policy π , the average cost $J_c(\pi)$ of (3.1) is obtained as a limit. From general results on MDPs there exists an optimal stationary policy for the average Lagrangian problem. Since the costs are positive, sample path arguments imply that this policy can be assumed non-idling. A standard Tauberian theorem [12] now implies that for each stationary non-idling policy, the average cost is the limit of the normalized discounted cost. Since (A1.a) holds in the discounted case (Lemma 4) where g_1, \ldots, g_L are strict priority policies, (A1.a) holds also for the average problem under the above conditions. Theorem 2 now implies the existence of an average cost optimal policy which randomizes between two strict priority policies.

Thus the result of Nain and Ross [19] extends to the Klimov problem, and this under both the discounted and the average cost criteria.

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